CO₂-gradient measurements using a parallel multi-analyzer setup

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Abstract
Accurate CO₂ concentration gradient measurements are needed for the computation of advective flux terms, which are part of the full Net Ecosystem Exchange (NEE) budget equation. A typical drawback of current gradient measurement designs in advection research is the inadequate sampling of complex flow phenomena using too few observation points in space and time. To overcome this drawback, a new measurement design is presented which allows the parallel measurement of several sampling points at a high frequency. Due to the multi-analyzer nature of the design, inter-instrument bias becomes more of a concern compared to conventional setups. Therefore a statistical approach is presented which allows for accurate observations of concentration gradients, which are typically small in relation to analyzer accuracy, to be obtained. The correction approach applies a conditional, time dependent signal correction. The correction depends on a mixing index based on cross correlation analysis, which characterizes the degree of mixing of the atmosphere between individual sample points. The approach assumes statistical properties of probability density functions (pdf) of concentration differences between a sample point and the field average which are common to the pdf's from several sample points. The validity of the assumptions made was successfully verified by Large Eddy Simulation (LES) using the model PALM. The study presents concentration time series before and after correction, measured at a 2 m height in the sub-canopy at the FLUXNET spruce forest site Waldstein-Weidenbrunnen (DE-Bay), analyzes the dependence of statistical parameters of pdf's from atmospheric parameters such as stratification, quantifies the errors and evaluates the performance of the bias correction approach. The improvements that are achieved by applying the bias correction approach are one order of magnitude larger than possible errors associated with it, which is a strong incentive to use the correction approach. In conclusion, the presented bias correction approach is well suited for – but not limited to – horizontal gradient measurements in a multi-analyzer setup, which would not have been reliable without this approach. Finally, possible future improvements of the bias correction approach are outlined and further fields of application indicated.
1 Introduction

Advection is a part of Net Ecosystem Exchange (NEE) of carbon dioxide, the determination of the latter being a primary focus of a world wide network of vegetation-atmosphere exchange measuring stations, the FLUXNET (Baldocchi et al., 2001). Not only are reliable measurements of advection lacking for most FLUXNET sites, but they continue to be a challenge even for specialized advection research experiments (e.g. Aubinet et al., 2003; Staebler and Fitzjarrald, 2004; Feigenwinter et al., 2008; Aubinet et al., 2010). Advection remains further to be a major reason for the night flux problem (Finnigan, 2008). Mathematically, scalar advection is the product of the mean spatial gradient of a scalar – CO₂ in the case of the current study – and the mean wind velocity, i.e. scalar transport with the mean flow. Advection is typically addressed as vertical advection (Lee, 1998; Baldocchi et al., 2000) and horizontal advection (Baldocchi et al., 2000; Aubinet et al., 2003).

There are two main conceptually different reasons why valid and representative advection measurements are difficult to obtain. One is the instrument related accuracy, with which scalar gradients and wind vectors of the mean flow can be measured. The other reason being undersampling of complex flow phenomena due to limited resources of real world experiments, thus yielding measurements which are not representative for a spatial (volume) and temporal (time period) mean but for a point only.

Vertical and horizontal advection pose different measurement challenges. With regards to vertical advection, reliable vertical CO₂ concentration gradients can be obtained due to vertical concentration gradients which are relatively large compared to sampling uncertainties. Measurements of vertical wind velocity are difficult to obtain, both for reasons of accuracy, precision, and resolution of sonic anemometers and particularly for reasons of the limited spatial representativity of a point measurement. Spatially representative measurements of vertical wind speed can never be obtained from a single point measurement in complex flow, due to theoretical reasons; therefore multi-tower measurements – possibly in combination with airborne measurements – are being suggested to improve spatial representativity of vertical wind measurements (e.g. Mahrt, 2010). Alternatively, the vertical wind velocity measurement problem is avoided by using a mass continuity approach, i.e. inferring vertical motion from horizontal divergence (e.g. Vickers and Mahrt, 2006; Montagnani et al., 2010) or a combination of the mass continuity approach and modeling (Canepa et al., 2010). Regarding horizontal advection, measurements of horizontal wind speed can be obtained with sufficiently high accuracy with sonic anemometers, even though they are often not spatially representative. Contrary, horizontal gradients are very difficult to measure with the required accuracy, because mean gradients are small in relation to instrument related uncertainty and difficult to measure at a large enough number of locations with a sufficiently high temporal resolution.

It is the main aim of this study to provide improvements for the measurement of horizontal CO₂ concentration gradients by means of a better temporal and potentially better spatial resolution. An improved resolution is needed for advection measurements in heterogeneous forests as could be shown by analyzing the effects of spatial heterogeneity and short lived phenomena on mean horizontal CO₂ concentration gradients (Siebicke et al., 2011).

The most commonly used setup for horizontal gradient measurements is based on a switching valve system (e.g. Burns et al., 2009), which uses a single closed-path infrared gas analyzer to sample several points one after the other (“sequential approach”), returning to the same sample point once every few minutes. There is an inherent tradeoff between achievable spatial and temporal resolution. The main benefit of this setup is a common analyzer for a number of sample locations, reducing the risk of bias between those points. The current study utilizes a multi-analyzer setup, featuring an individual closed-path infrared gas analyzer for every measurement point, enabling simultaneous measurements of all points (“parallel approach”) with a high frequency. Temporal resolution is no longer parasitic to spatial resolution, the latter depending on available resources only. With ten individual analyzers used, the spatial resolution is on the order of a sequential system. Thus the system described is capable...
of making measurements which are representative in the temporal domain since it can observe all relevant temporal scales of the CO₂ concentration signal. Valid concentration measurements need to be both precise and accurate. Precision of the parallel approach used in this study is higher compared to the conventional sequential approach because there are potentially much more values available in one averaging interval, thus reducing random error. The advance in the number of values is proportional to the number of sample locations per analyzer for the sequential approach. Lower accuracy of a multi-analyzer setup compared to a single analyzer setup due to inter-instrument bias is the major drawback of the parallel approach, in addition to higher resource requirements. Bias can be reduced by careful system design and frequent calibration against accurate, known standards. Section 2.2 lists technical measures that have been taken to that end for the presented system. How to deal with the remaining bias will be the topic of the rest of the paper. The basic assumption regarding concentration differences originating from natural gradients stated in Sect. 2.4.2, which is the justification of the proposed bias correction approach, has been implicitly used by Aubinet et al. (2003) and applied for time series correction in a simple, time independent manner whereas the current study applies a conditional, time dependent signal correction. Previous studies using more than one closed path gas analyzer in a multiplexer system with multiple sampling inlets have often used co-located inlets to deal with time dependent inter-instrument bias (e.g. Sun et al., 2007), and the same procedure was applied to vertical profile measurements at the site of the current study. However, due to the characteristics of the multi-analyzer system presented in this study with only one inlet per analyzer, co-located inlets cannot be used in the same way and a new approach is needed.

It should be noted that the term “CO₂ concentration” is used throughout this paper to describe basic principles in a consistent way. It specifically refers to “molar fraction” in units of mol mol⁻¹ or µmol mol⁻¹, which were used for all measured values presented herein, whereas it refers to “CO₂ density” in units of kg m⁻³ for modelled values from the Large Eddy Simulation study (Sect. 2.5 and Sect. 3). However, further applications of the ideas about bias correction presented in this paper may prefer to describe CO₂ in terms of “mixing ratio” in units of kg kg⁻¹ (Kowalski and Serrano-Ortiz, 2007).

2 Material and methods

2.1 Site

Measurements were carried out at the FLUXNET site Waldstein-Weidenbrunnen (DE-Bay), 50° 08′ 31″ N, 11° 52′ 01″ E, a hill site in the Fichtelgebirge Mountains in Southern Germany. The 25 m high spruce stand is on the upper section of a hill, 775 m a.s.l., with a 3° slope facing south-west. The site is described in detail in Gerstberger et al. (2004) and a summary of background data can be found in Staudt and Foken (2007).

2.2 Instrumental setup

Wind vector and CO₂ concentration time series were recorded along horizontal transects at a 2.25 m height in the sub-canopy space. The spatial setup of sub-canopy sample locations is shown in Fig. 1. Ten CO₂ concentration sample points were distributed between an along slope transect from north-east to south-west (5 sample points) and an across slope transect from north-west to south-east (6 sample points), including one common point. Each point was sampled by an individual closed-path infrared gas analyzer. Instruments used were five LI-6262, one LI-6251 (LI-COR Biosciences Inc.), four BINOS (Leybold Heraeus GmbH). In addition to CO₂ concentration measurements at a 2.25 m height, sample locations M5, M6, M7, M8, M9, M10 (see Fig. 1) were equipped with sonic anemometers (USA-1, METEK GmbH) to measure wind speed, wind direction and sonic temperature at the same height. CO₂ concentration measurements are available with a frequency of 1 Hz at each sample point, sonic data were recorded at a 20 Hz frequency. To reduce the risk of systematic differences between individual closed-path gas analyzers the system was carefully designed to
Avoid any possible bias of the concentration measurement from differences in pressure or temperature (sample air temperature, ambient analyzer temperature, radiation). All CO₂ closed-path gas analyzers shared a common housing in a central position with controlled conditions resulting in a constant common temperature and common pressure regime. Moreover, all analyzers shared a common tailor-made automatic calibration system, using high precision reference gases (accuracy 0.1 µmol mol⁻¹). The calibration routine included an automatic calibration every 4 h using two reference concentrations. In addition to factory calibration, each instrument’s polynomial calibration function was established on site, using multiple standards. The polynomial was checked before and during the experiment.

Individual technical measures taken to avoid systematic inter-instrument bias include the following:

- The length of tubing connecting each sample point with the corresponding gas analyzer was exactly 75 m for every point. Sample tubes used were of polyethylene-aluminum composite structure, model DEKABON 1300-M060X (Serto AG, Fulda-brück, Germany) with an inner diameter of 4 mm.

- Large diameter line intake air filters were checked regularly and replaced synchronously at all points, if necessary.

- Common ambient temperature and pressure for all gas analyzers and calibration unit, including radiation protection, active automatic temperature control by heating and cooling as well as carefully designed ambient air circulation.

- Quality control of performance of automatic temperature control system, making sure that ambient air temperatures measured at several points surrounding the gas analyzers remain within acceptable range.

- Temperature adaptation for sample lines, to allow the temperature of sample air in all sample lines to equilibrate to a common temperature prior to entering the analyzer.

- Common temperature and radiation shielding for reference gases.

- Minimization of dead volumes in calibration and valve system to ensure turbulent flow conditions and avoid contamination by previous samples.

- Flow rate of 2 L min⁻¹ (Reynolds number Re = 2520) above critical flow rate of 1.8 L min⁻¹ at critical Reynolds number (Re crit = 2320) to ensure turbulent flow conditions in all tubes, at the same time keeping the flow rate as low as possible to minimize pressure drop across the system.

- Regular flow rate check and adjustment for all sample lines.

- Bypass system to avoid back pressure effects during calibration, featuring a low pressure drop bypass flow rate control device to ensure minimum necessary bypass flow and avoid possible reverse flow and sample contamination by ambient air.

- One common pump downstream of the analyzers to reduce effects of the pump on the concentration signals and to guarantee common pressure for all analyzers, assuming equal pipe geometry of all sample lines.

- Automatic control of constant overall system flow rate by mass flow controller.

- Passive system to allow for pressure equilibration between sample cells of individual gas analyzers by connecting all analyzer outlets to a manifold with a sufficiently large diameter.

- Pre-assembly measurement and evaluation of the pressure drop caused by individual system components to ensure that associated errors of the CO₂ concentration measurements are below accepted threshold.

- Vacuum and over pressure assisted leak check for the complete system to rule out sample contamination by ambient air.
2.3 Data set

The data set was collected during the second intensive observation period (IOP2), 1 June to 15 July 2008 of the EGER (“ExchanGE processes in mountainous Regions”) experiment (Foken et al., 2011). 24.6 days worth of data were used for the analysis, i.e. 1181 half hourly values taken from a window of 32.0 days (11 June to 13 July). Periods were excluded from the analysis when instruments were powered off or obviously malfunctioning.

2.4 Theoretical considerations regarding concentration differences

2.4.1 Bias

An observed concentration difference between two spatially separated sample points is the sum of a concentration difference originating from a natural atmospheric concentration gradient and the inter instrument bias, the latter being determined by systematic (bias) and random error of the individual instruments. We will refer to this composite concentration difference also as a concentration offset, $\Delta c$. While random error of the instruments is a minor concern in the current study due to sufficiently long averaging period, instrument bias can be reduced by calibration against known standards. The calibration procedure used in this study was outlined in Sect. 2.2. The remaining bias is the sum of the error of the calibration plus the instrument drift between two consecutive calibration events. This remaining bias cannot be removed by calibration since it is intrinsic to the calibration procedure itself. However, a statistical approach detailed in Sect. 2.7 can help to distinguish between remaining bias and concentration differences originating from natural gradients based on the observed signal.

2.4.2 Natural concentration differences

To separate concentration differences originating from natural gradients between two spatially disjunct (i.e. up to a few tens of meters) sample points from instrument bias the following assumption is made and is the basis for bias correction used in the current study: for certain meteorological conditions the concentration time series observed simultaneously at the two locations can be statistically linked to a reference concentration which is common to both sample locations. To be more precise, under the condition of well mixed, i.e. sufficiently turbulent atmospheric conditions (hereafter “mixed” conditions) the concentration difference between the two locations which is most likely to be observed is zero. If this statement is true for the concentration difference between any two points, it can also be applied to the difference between the concentration at one sample location $c_i$, and the spatial average concentration of the sample point field $\bar{c}(t)$ at a given time $t$. $\bar{c}(t)$, which serves as a reference concentration, describes the background concentration of the sample point field at $t$ using the median field concentration according to Eq. (1)

$$\bar{c} = \begin{cases} 
  c_{n+1} & n \text{ odd} \\
  \frac{1}{2} (c_n + c_{n+1}) & n \text{ even} 
\end{cases}$$

with $n$ observations ($c_1(t), c_2(t), \ldots, c_n(t)$).

The statistical measure describing the concentration difference most likely to be observed is the mode of the probability density distribution (pdf) of the concentration differences $c_i(t) - \bar{c}(t)$, which is assumed to be close to zero under the condition of well mixed i.e. sufficiently turbulent atmospheric conditions. This is illustrated in Fig. 2b for two hypothetical time series $c_1(t) = 7.6, 5.8, 5, 4.6, 5.6$ and $c_2(t) = 7.6, 7.5, 3.5, 4.5, 6.5$, displayed in Fig. 2a. The characteristics of turbulence justify the assumed mode of the pdf to be close to zero, i.e. turbulence consists of temporal perturbations of a mean state which are stochastic and relatively short in duration compared to the observation period. The mode is zero even though the time series $c_1(t)$ and $c_2(t)$ given in Fig. 2a have a different mean (temporal mean indicated by overline): $\bar{c}_1(t) = 5.7$ and $\bar{c}_2(t) = 5.3$, and even though the mean of the concentration difference $c_i(t) - \bar{c}(t)$ is different from zero: $\bar{c}_1(t) - \bar{c}(t) = 0.2$ and $\bar{c}_2(t) - \bar{c}(t) = -1.3$. Yet, the mode of the concentration differences $c_i(t) - \bar{c}(t)$ is close to zero and hence the assumption of zero concentration difference is justified.
c_2(t) - \tilde{c}(t) = -0.2.

For atmospheric conditions without turbulent mixing (hereafter "non mixed" conditions) above stated assumption does not need to be fulfilled. Since there is no effective mechanism of mixing, two sample locations can be continuously exposed to air masses with different concentrations – see concentration time series c_1(t) = 4,3,2,5,2,1,3,2,3 and c_2(t) = 8,7,6,4,6,5,6,7,6 in Fig. 2b – i.e. there is a persistent natural gradient and no common background concentration is observed at both sample points. Thus, the two points will most frequently sample a concentration difference which represents this gradient, and the mode of the probability density distribution is non zero, Fig. 2d.

All combinations of the well mixed and non mixed case are possible. It depends on turbulence statistics and the length of the time series incorporated in the probability density distribution whether mixing is sufficient to produce a mode of the pdf close to zero or not. A method to quantify the degree of mixing will be presented in Sect. 2.6.

2.5 Large Eddy Simulation

Large Eddy Simulation (LES) is a tool that is used to study turbulence related processes in the atmospheric boundary layer. It can therefore be used to extract statistical properties of turbulence for the well mixed case. A LES study was performed to test whether the above stated assumption is true, i.e. whether the mode of the pdf of the concentration differences c_i(t) - \tilde{c}(t) between one sample location and the spatial average of the sample point field is close to zero for well mixed conditions even though there is a vertically non homogeneous source-sink distribution and a mean spatial concentration gradient \delta c/\delta y with concentration c and horizontal distance y. The simulation does not intend to perfectly mimic subcanopy conditions but to test general statistical properties of turbulent mixing, i.e. whether strong turbulent mixing is able to allow the average field background concentration \tilde{c}(t) to emerge as the dominant mode of the pdf rather than local sources or sinks producing the dominant mode.

The LES model used in this study is the Parallelised LES Model (PALM) that has been developed at the Institute of Meteorology at the Leibniz University in Hannover, Germany. Detailed information on the LES approach, model equations and numerical schemes applied in PALM are given in Raasch and Schrötter (2001) or – continuously updated – on-line on the homepage of the PALM group (Raasch, 2010). Two separate simulation configurations were chosen: "case A" was initiated with a horizontal concentration gradient, "case B" is characterized by a horizontal source strength gradient. The two configurations will be described in the following paragraphs.

For this study, we used PALM with an additional prognostic equation for a scalar quantity, so that the temporal development of scalar concentration c evolving from a horizontally homogeneous (case A) or horizontally heterogeneous (case B), continuous volume source, extending throughout the xy-cross section of the model domain at a 2.5 m height, and from a sink, extending throughout the xy-cross section of the model domain at a 27.5 m height, could be simulated. The scalar concentration was initialized everywhere in the model domain with a basic concentration of 6.997 x 10^{-4} kg m^{-3}, equivalent to a 378 µmol mol^{-1} CO_2 background concentration observed at the site.

For case A, an additional concentration with a horizontal gradient of scalar concentration was prescribed along the y-direction on top of the basic concentration (Fig. 3b, see left axis). This gradient \delta c/\delta y was 3.038 x 10^{-7} kg m^{-4} for y \leq \frac{L_y}{2}, while it was -3.038 x 10^{-7} kg m^{-4} for y > \frac{L_y}{2}, where L_y is the extension of the model domain along the y-direction. This prescribed gradient is equivalent to 0.16 µmol mol^{-1} m^{-1} which deliberately has been chosen to represent the maximum of gradients observed in the field at the site under study and published for other sites (Aubinet et al., 2003; Heinesch et al., 2007) during stable stratification, even though gradients are smaller during neutral and unstable stratification, i.e. the stratification regime present in the LES. Therefore, the LES with strong gradients tests a worst case scenario. At y = 0 and y = L_y this additional contribution to the total concentration was 0. The scalar quantity was released at a 2.5 m height with a source strength of 8.8 x 10^{-8} kg m^{-3} s^{-1}, while
the sink at a 27.5 m height had a strength of $-8.8 \times 10^{-8} \text{kg m}^{-3} \text{s}^{-1}$ (approx. equivalent to maximum daytime Net Ecosystem Exchange of $-20 \mu\text{mol m}^{-2} \text{s}^{-1}$ observed at the site).

For case B no additional concentration gradient was prescribed but a horizontal gradient in the source strength $s$ along the $y$-direction was introduced (Fig. 3b, see right axis). With a basic source strength of $4.4 \times 10^{-8} \text{kg m}^{-3} \text{s}^{-1}$ at $y = 0$ and $y = L_y$, the source strength gradient $\frac{\partial s}{\partial y}$ was $5.5 \times 10^{-11} \text{kg m}^{-4} \text{ s}^{-1}$ for $y \leq \frac{L_y}{2}$, while it was $-5.5 \times 10^{-11} \text{kg m}^{-3} \text{s}^{-1}$ for $y > \frac{L_y}{2}$. Case B has a mean source strength of $8.8 \times 10^{-8} \text{kg m}^{-3} \text{s}^{-1}$, which is equivalent to the constant source strength in case A (again approx. equivalent to maximum daytime Net Ecosystem Exchange of $-20 \mu\text{mol m}^{-2} \text{s}^{-1}$ observed at the site), with 50% of that value at the domain borders $y = 0$ and $y = L_y$ and 150% of that value in the center of the domain at $\frac{L_y}{2}$. Source and sink height and sink strength of case B are equivalent to case A.

The release and the extraction of the scalar quantity started after a spin-up time of 2 h. The simulations were initialized with wind profiles obtained from a one-dimensional pruned. The geostrophic wind $(u_g, v_g)$ was prescribed as $(3 \text{ m s}^{-1}, 0 \text{ m s}^{-1})$ while $u$ and $v$ correspond to the $x$- and $y$-direction, respectively. The roughness length was 0.1 m. Initially, the potential temperature was constant up to a height of 400 m. At larger heights the potential temperature increased by 0.01 K m$^{-1}$. At the bottom boundary of the model domain a near-surface heat flux of 0.01 K m$^{-1}$ was prescribed, so that a convective boundary layer developed with time. The Coriolis force was taken into account in the simulation and the Coriolis parameters used were that obtained for a geographical latitude of 55°. Time series of scalar concentration were recorded at 16 observation points within the $xy$-cross section of the model domain (3.2 x 3.2 x 2 km, 640 x 640 x 256 grid points, grid stretching in the vertical direction above 1000 m with a maximum grid size of 20 m) at a height of 17.5 m beginning from the first release of scalar quantity until the end of the LES 7200 s later. Figure 3 shows the location of the virtual observation points. The coordinates of the 16 observation points were composed out of the $x$-coordinates (760 m, 785 m, 810 m, 835 m) and $y$-coordinates (760 m, 785 m, 810 m, 835 m). Thus, the distance between two adjacent observation points along the $x$- or $y$-direction was 25 m.

2.6 Mixing index

A “mixing index” $MI$ was formulated to quantify the degree of mixing between the real world sample points given in Fig. 1. A threshold value $MI_c$ was then used to separate conditions which satisfy the assumption from those violating it. The mixing index $MI$ is based on the cross correlation $R_{c_1c_2}(t)$ of the simultaneous concentration time series $c_1(t)$ and $c_2(t)$ of spatially separated sample locations normalized by their mean variance $\sigma^2$. The cross correlation function is given as

$$R_{c_1c_2}(t) = \frac{1}{T_F} \int_{-T_F/2}^{T_F/2} c_1(t) \cdot c_2(t + \tau) d\tau$$

with time lag $\tau$ between concentration time series $c_1(t)$ and $c_2(t)$, $T_F$ being the length of the time window of $c_1(t)$ and $c_2(t)$ and $\tau \in [-T_F, T_F]$. $MI$ then writes:

$$MI = \max (|R_{c_1c_2}(t)|) \cdot \left( \frac{\sigma^2_{c_1} + \sigma^2_{c_2}}{2} \right)^{-1}$$

More specifically, $MI$ was calculated using the mean cross correlation of CO$_2$ concentration time series $c_5$ and $c_6$ recorded at a sample point pair oriented along the terrain slope (locations M5, M6) and $c_8$ and $c_9$ recorded at a sample point pair oriented across the slope (M5, M8) divided by the mean field variance of all concentration time series $c_5$, $c_6$, ..., $c_{14}$ at sample locations M5, M6, ..., M14 using a window length of $T_F = 60$ min.

The critical mixing index $MI_c$ was empirically inferred from the density distribution of $MI$ given in Fig. 4a. Sensible values were found to be in the range $MI_c \in [0.06, 0.12]$,
corresponding to a sharp bend in the density distribution separating $M_I$'s representative of well mixed daytime conditions (distribution tail to the right of $M_{IC}$ in Fig. 4a) from low $M_I$'s representative of night time conditions with little mixing (distribution peak to the left of $M_{IC}$ in Fig. 4a). Figure 4b presents a typical diurnal cycle of the mixing index which is clearly separated into mixed and non mixed conditions by $M_{IC}$.

### 2.7 Bias correction

Instrument related bias of the CO$_2$ concentration signal was observed to vary over time. It is therefore appropriate to apply a bias correction that is time dependent, too. Analyzer specific values of instrument bias $\Delta c_i$ were computed for every 60-minute interval $T_i$ of the concentration time series $c_i(t)$ by finding the mode ($\text{max}$(density)) of the probability density distribution (pdf) of the instantaneous concentration differences of an individual analyzer $c_i(t)$ relative to the field average concentration $\bar{c}(t)$ according to

$$\Delta c_i = \text{max}(pdf(c_i(t) - \bar{c}(t))) \quad (4)$$

with $\bar{c}(t)$ defined in Eq. (1) and the total number of analyzers $n = 10$. Identifying the mode of the pdf requires a robust estimate of the distribution. A comparison of histogram based and kernel-density-estimator based approaches showed that the latter are superior in terms of robustness relative to scatter in the distribution, which is a valuable feature particularly for limited sample sizes. Density estimates were generated using a moving window Gaussian kernel for smoothing (Wand and Jones, 1995). The optimal width of the window was adaptively and automatically found using pilot-density-estimates according to Sheather and Jones (1991), implemented in the `bkde` function of the `KernSmooth` library (Ripley, 2009) provided with R (R Development Core Team, 2009), also providing the `bkde` function which was used to estimate the density. Having found an individual bias value for every analyzer, the mixing index was checked to decide whether concentration time series correction was applicable. For well mixed conditions, i.e. $M_I \geq M_{IC}$, the observed 60-minute concentration time series $c_i(t)$ of every analyzer was shifted by the analyzer specific bias value $\Delta c_i$ found for the given 60-minute interval, yielding the bias corrected concentration time series $c_{i,\text{corr}}(t)$ according to Eq. (5).

$$c_{i,\text{corr}}(t) = c_i(t) - \Delta c_i \quad \text{for} \quad M_I \geq M_{IC} \quad (5)$$

For $M_I < M_{IC}$ the correction was applied using the last valid bias value satisfying $M_I \geq M_{IC}$.

In order to verify that concentration offsets $\Delta c_i$ found are related to slow drift of the analyzers (instrument bias) rather than driven by meteorological forcing of natural concentration gradients, a regression analysis was performed studying the correlation of $\Delta c_i$ versus ambient air temperature, pressure and atmospheric stability $\zeta$, respectively. The stability parameter $\zeta$ is defined as $\zeta = (z - d)/L$ with measurement height $z$, displacement height $d$ and Obukhov-length $L$. No significant correlation was found between the concentration offset and the three meteorological parameters, which is an indication that the calculated offset $\Delta c_i$ is dominated by instrument bias and should therefore be removed with the proposed conditional bias correction approach, respecting $M_I \geq M_{IC}$.

Because, even under mixed conditions, natural concentration differences could account for a (very small) part of the observed concentration offsets $\Delta c_i$, an error analysis was performed. The aim was to quantify the benefit of the application of the bias correction approach in a hypothetical "worst case" scenario, i.e. assuming that observed concentration offsets $\Delta c_i$ are solely determined by natural concentration differences rather than instrument bias. A relative error is defined in Eq. (6), describing the ratio of the error possibly attributed to the bias correction approach to the improvement achieved by the correction, which can be expressed as the span of the range of instrument bias ("drift span"). This relative error writes

$$\text{error}_{rel} = \frac{Q_1(\Delta \text{off}_i) - Q_4(\Delta \text{off}_i)}{\text{max}(\text{off}_i) - \text{min}(\text{off}_i)} \quad (6)$$
with the change of the concentration offset $\Delta c_j$ between two consecutive 60-minute intervals $\Delta \text{off}_j = \Delta c_j(t) - \Delta c_j(t-60 \text{ min})$ and with $Q_1$ and $Q_2$ being the 25% and 75% quartiles of the density distribution, respectively, which reflect a typical range of $\Delta \text{off}_j$.

### 2.8 Horizontal advection

The fact that density distributions of concentration differences can have a mode of zero and a non zero mean, as seen in Fig. 5b, is a feature with importance for the computation of advection, because only a non zero mean gradient $\frac{\partial c}{\partial x} \neq 0$ and/or $\frac{\partial c}{\partial y} \neq 0$ can generate a non zero horizontal advection term $F_{HA}$

$$F_{HA} = \frac{1}{V_m} \int_0^h \left( \bar{u}(z) \frac{\partial \bar{c}}{\partial x} + \bar{v}(z) \frac{\partial \bar{c}}{\partial y} \right) dz \quad (7)$$

with the molar volume of dry air $V_m$, CO$_2$ concentration $c$, horizontal distances $x$ and $y$, measurement height $z$ above ground, horizontal wind velocity $u$ along the $x$-direction and horizontal wind velocity $v$ along the $y$-direction.

### 3 Results

After presenting results of the LES study, which contribute to the acceptance of the assumptions stated in Sect. 2.4.2, this section subsequently presents results of measured CO$_2$ concentration time series and gradients before and after applying the conditional bias correction as well as statistics about the improvement which can be achieved by the correction. Furthermore, observed concentration differences are put in the context of atmospheric stratification.

The results of the LES study demonstrate that for the given simulation the assumption stated in Sect. 2.4.2 is valid, i.e. the mode of the density distribution of the concentration difference between any sample point and the sample point field average is essentially zero, Fig. 5a. Since both case A and case B lead to the same conclusion, only data of case B are shown in Fig. 5. Observed deviations of the density distribution mode from zero are insignificant, with the maximum deviation, considering all instrument’s distributions, divided by the mean distance of the sample point from the sample point field center, accounting for a 2.0% fraction only of the prescribed concentration gradient in the LES (case A). For case B the maximum deviations of the mode from zero were $+0.015$ and $-0.025$ µmol mol$^{-1}$. Dividing this range of distribution modes by the range of the distributions means yields a fraction of 0.15. Considering the small gradients under well mixed conditions, this is a very small error. Conditions with large gradients are not an issue because they are excluded by the mixing index filter and are not used to determine concentration offsets when applying the bias correction approach.

The given deviations of the pdf’s modes translate to an error attributed to estimates of the horizontal advective flux component, if estimates are based on concentration measurements corrected using the bias correction approach and thus removing the small deviation of the mode from zero. This potential error in the advection estimate is small compared to other uncertainties typically associated with advection estimates, e.g. due to an insufficient number of sampling points in space such as the often limited number of observation height levels of horizontal gradients.

An important feature of the density distributions shown is their skewness, separating mode and mean of a given distribution as illustrated in Fig. 5b for two selected sample points. The difference in the mean values of the density distributions is due to the concentration gradient and source-sink distribution prescribed in the LES. It thus demonstrates that it is possible to compute advective flux terms even from distributions with mode equal to zero, since the mean gradient, which is necessary to compute $F_{HA}$ according to Eq. (7), is expressed in the mean which does not need to be zero even though the mode is essentially zero.

In order to evaluate the performance of the bias correction, Fig. 6a shows the CO$_2$ concentration evolution during one day measured at ten locations in the sub-canopy on
29 June 2008 without bias correction but including calibration using known reference gas standards. Figure 6b presents the same data after applying the bias correction. The comparison of the two figures clearly demonstrates that the bias correction is able to remove systematic concentration offsets between different analyzers in the uncorrected measurements (Fig. 6a). The offsets are most obvious during well mixed daytime conditions – when natural concentration differences are relatively small – and could be eliminated successfully in the bias corrected time series at all times of the day (Fig. 6b).

Inter-instrument bias leads to relatively constant offsets between individual concentration measurements \( \Delta c(t) \) during daytime conditions (Fig. 6a), exactly matching the period of a high mixing index (Fig. 4b). The minor importance of concentration differences due to natural gradients during well mixed conditions is the reason why inter-instrument bias becomes the prominent component of observed inter-instrument concentration differences (compare also Fig. 9 and Fig. 10). Well mixed conditions with \( MI \geq MI_c \) and \( MI_c = 0.13 \) were observed every day during the experiment, accounting for 30% of the entire data set. There are a few cases where mixed conditions are present for short isolated periods (e.g. one or two 60-minute \( MI \) values) at transition times in the early morning or sometimes in the early evening.

While Fig. 6a and 6b presented \( \text{CO}_2 \) concentration time series before and after bias correction on 29 June 2008, Fig. 7 displays an example of corresponding density distributions of concentration differences during a well mixed 60-minute period at midday of the same day, which were used during bias correction. Probability density distributions with analyzer specific non-zero distribution modes in the uncorrected data of Fig. 7a have been shifted by their mode so that the new mode of the distributions is equal to zero after bias correction (Fig. 7b). Figure 7b also emphasizes sample location specific differences of the distribution shape, such as different skewness and kurtosis, which is an effect of natural concentration gradients being unique for every sample location.

Having discussed probability density distributions above for an ideal case with mixed conditions, Fig. 8 demonstrates the effect of atmospheric stratification \( (\zeta) \) and the degree of mixing \( (MI) \) on the shape of selected 60-minute probability density distributions of concentration differences, which mark typical conditions during the course of a fair weather day, 29 June 2008. Distributions of the well mixed case in Fig. 8c are unimodal and show high kurtosis. This is beneficial for the reliable estimation of the mode, which is necessary for bias correction. High kurtosis is a consequence of small natural horizontal and vertical gradients during well mixed conditions in the middle of the day. Figures 8b and 8d represent transition periods between night and day and between day and night, respectively, while Fig. 8a and 8e are examples of night time conditions, with Fig. 8e being a representative example for conditions with katabatic sub-canopy drainage flow under very stable conditions. The kurtosis of the distributions correlates with \( \zeta \) (indicator for atmospheric stratification) as well as with \( MI \) (indicator for turbulent mixing), the result being that kurtosis decreases and skewness often increases with increasing stability parameter \( \zeta \) and decreasing mixing index \( MI \). This is due to large horizontal and vertical scalar concentration gradients during such conditions, also potentially causing multimodal distributions (Fig. 8b and 8d), which can lead to disambiguities concerning the relevant mode if they were to be used for bias correction, which they are not due to the mixing index condition. However, the effect of atmospheric stability \( \zeta \) is not uniform, meaning that multiple modes and skewed distributions (Fig. 8b) and low kurtosis (Fig. 8d) are more pronounced during transition periods with moderate vertical exchange, whereas the night time cases such as Fig. 8e with the highest stability parameter \( \zeta \) and least vertical exchange are less skewed and more homogeneous with respect to different sample locations. The absence of vertical exchange results in horizontally relatively homogeneous sub-canopy scalar concentrations even though there are large vertical gradients.

Figures 9 and 10 demonstrate that observed concentration offsets \( \Delta c(t) \) can be separated into offsets which are mainly determined by instrument bias alone and into offsets which are determined by instrument bias as well as significant natural concentration differences. Figure 9a shows offset time series over two days with a succession of mixed daytime conditions (approx. 8 h to 16 h) with little scatter in the offset time series when
natural gradients are small and offsets are mainly controlled by low frequency instrument bias and night time conditions with high scatter and large absolute values in the offsets time series when natural gradients are the predominant cause. After applying the mixing index to filter the offset time series, those periods with predominant natural gradients were effectively excluded (Fig. 9a). The remaining offsets are controlled by instrument bias and can therefore be used in the bias correction approach.

The different offset characteristics during daytime and nighttime described above are due to the dependence of natural concentration differences on the mixing index and atmospheric stability, both of which have a distinct daily cycle. Figure 10 illustrates the dependence of concentration offsets on the mixing index $MI$. For low values of $MI$, natural horizontal gradients are large, as a result of horizontal source heterogeneities and potential mixing of a vertical concentration profile with large vertical gradients. For larger values of $MI$, particularly for $MI > MI_c$, offsets are relatively small. Figure 10 displays the dependence of only natural concentration differences on $MI$. The graph shows offsets with instrument bias removed by subtracting a low frequency component (see figure caption for details). The fact that this technique does not perfectly separate natural concentration differences from instrument bias explains the scatter and outliers in Fig. 10 which are present even at higher values of $MI$. The majority of data points (indicated by solid lines for the 25% and 75% quantiles) in Fig. 10 is close to zero concentration difference for higher values of $MI$ ($MI > MI_c$). That indicates that there are no major natural concentration differences under those conditions which could unintentionally be removed by the bias correction approach.

To compare the benefits of the bias correction approach with potential errors, Table 1 displays results of an error analysis, listing the potential for improvement by using the bias correction approach (drift span), an estimate of the potential absolute error $(Q_{i}(\Delta off_i) - Q_{i}(\Delta off_i))$ and the relative error $(error_{rel})$ for ten sampling locations. Values of the relative error are on the order of 10%, which is a satisfying result, keeping in mind that those are “worst case” values pretending that offsets during mixed conditions, i.e. when the bias correction is applied, were purely caused by natural gradients, which they are not in reality. Therefore the true relative error will be even smaller than values given in Table 1 for $error_{rel}$.

4 Discussion

There are three issues connected to the quality of the bias correction approach which shall be discussed in this section: potential underestimation of natural concentration differences (signal loss), tradeoff between limiting instrument drift and limiting signal loss, and finding the appropriate window length when applying the bias correction.

The previous section has shown that the improvements that were achieved by applying the bias correction approach are one order of magnitude larger than possible errors associated with it, which is a strong incentive to use the correction approach. However, there is potential for losing part of the natural concentration gradients when applying the correction, due to possibly imperfect separation of instrument bias and concentration differences originating from a natural gradient, even during mixed conditions. A quantification of this phenomenon was given in Table 1. The acceptance of this relatively small potential error when applying the bias correction approach needs to be compared to errors which are likely to be attributed to the gradient measurements with no correction applied. It is well known from various advection experiments that instrument related bias between sampling points can easily be larger than the natural horizontal concentration gradients which are to be observed, the latter often being small compared to the accuracy of the measurement. This in turn can lead to considerable overestimation of the absolute value of horizontal advection, which is one of the reasons why including the horizontal advection flux term in the Net Ecosystem Exchange (NEE) budget equation often leads to increased scatter of NEE and does not necessarily produce reliable NEE estimates. As a consequence, NEE is often computed using the turbulent and storage fluxes only. We suggest that rather than including a noisy and potentially too large advection estimate in the NEE equation, it is better to include a bias corrected estimate of horizontal advection. Doing so and at the
same time accounting for vertical advection – the same arguments apply here as to avoid overestimation and noise – should give more realistic NEE estimates than those obtained from turbulent and storage flux alone.

When applying the bias correction, a balance should be found between limiting the effect of instrument drift on the gradient measurements and signal loss by potential underestimation of natural gradients. This balance can be tuned by the choice of the value for the critical mixing index $MI_c$. A high value of $MI_c$ better preserves natural gradients because bias correction values are only determined from data during well mixed conditions and therefore can not eliminate natural gradients during other conditions, particularly at night when natural gradients are typically large. A low value of $MI_c$ removes instrument drift more thoroughly since bias correction values can be found more often, i.e. from well mixed as well as partly mixed conditions. Therefore, we recommend to choose a higher $MI_c$ the more stable the analyzer is and just low enough to allow the instrument to "survive" periods during which no bias correction values can be found (i.e. nighttime) using previously established correction values (inherited from daytime) without facing prohibitive instrument drift during those periods.

The third issue is finding the appropriate window length $T_F$ when applying the bias correction. This is the length of the time series used to compute density distributions of concentration differences (pdf) to find their mode as outlined in Sect. 2.7. For this study the window length was chosen to be $T_F = 60$ min. The higher the instrument drift is, the shorter this window has to be in order to find a mode which is representative for the instrument bias during that time window and not affected by a significant trend of the bias. On the other hand, choosing the window as long as possible helps to preserve natural gradients which are persistent for longer periods, since persistent natural gradients with periods longer than the window length and present during non mixed conditions, and therefore affecting the mode of the pdf, are removed by the bias correction for $MI > MI_c$. However, we can conclude from the data that it is not satisfactory to choose an infinite window length (such as the time constant bias correction applied by Aubinet et al., 2003) in order to preserve natural gradients because observed instrument bias is subject to drift over time. Given the window length of 60-minutes used in the current study, the concentration difference error due to signal loss of natural concentration differences during the day has been shown to be smaller than the error of the concentration offset which would be caused by the drift of the instrument bias if the latter was corrected by a time constant bias value. Future studies should test window lengths larger than 60-minutes, particularly when using more stable analyzers.

Future work on the improvement of the bias correction approach should include a refined condition to test which data should be used when determining the pdf and the bias. Rather than using fixed 60-minute intervals to determine $MI$ and accepting all data in a 60-minute interval satisfying $MI > MI_c$, a more fine grained selection of data entering the pdf can be used to select only those parts of the time series which have common properties at more than one sample point for a time period on the order of the duration of coherent structures, i.e. seconds to minutes. Among the tools which can be used to find common properties within the time series are cross correlation analysis and pattern recognition. Thereby only data with similar concentration at several sample points will enter the pdf. This helps to exclude the influence of natural gradients on the mode of the pdf, which will then be determined by instrument bias alone. Such short term correlation of time series at several sample points by tracking individual structures in the time series should be done for sample point pairs rather than using properties of the complete sample point field. These pair wise correlations then need to be linked together by choosing different configurations of sample point pairs and combining their information.

Future work can also test the applicability of the bias correction approach to sensor networks with a possibly large number of sampling points. The approach can be used when working with sensors which have a relatively high resolution but suffer from low accuracy. Whereas those sensors will deliver the fine structure (high frequency part) of the time series, the bias correction approach corrects constant and drifting instrument bias (low frequency part) and thus ensures the accuracy of the measurements.
5 Conclusions

This paper has presented a measurement design capable of addressing the issue of inadequate sampling of natural concentration gradients in the temporal domain – a common characteristic of many advection measurement setups – by increasing the temporal resolution of the gradient measurements. Observing gradients with a sufficiently high temporal resolution and therefore capturing as much information as possible over a large range of temporal scales is crucial for reliable advection estimates computed from concentration gradients. In order to produce accurate gradient measurements in a multi-analyzer setup, an approach was presented which adequately addresses the problem of inter-instrument bias. It was shown that the uncertainties associated with this approach are one order of magnitude smaller compared to the benefit achieved. The proposed bias correction approach is therefore a suitable tool at least for multi-analyzer setups measuring horizontal gradients at one height, given a certain proximity of individual sampling locations. There might also be benefits from applying the bias correction approach to sequentially measured data from switching valve systems in a single-analyzer setup. It should be tested in the future whether the bias correction approach can be transferred to measurements of vertical gradients, although care has to be taken due to strong systematic vertical gradients particularly at night in the case of CO\textsubscript{2} concentration. The concept outlined in the current paper is not limited to measurements of CO\textsubscript{2} concentration but useful for the accurate observation of gradients of other scalars, too. Furthermore, it is not limited to gradient measurements for the computation of advective flux components but is worth considering for any gradient based flux measurement application. Finally, the bias correction approach is useful for the relative adjustment of signal levels between individual sensors in any kind of sensor network that samples phenomena which – at least part of the time – lead to common characteristics of the observed signal at several locations in the network.

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Table 1. Offset statistics and error analysis for ten sample points, i.e. ten analyzers, demonstrating the correction potential of the bias correction approach ("drift span"), typical values for the maximum error possibly attributed to the bias correction approach for a "worst case" scenario (from quartile Q₁(Δoff₁) to quartile Q₄(Δoff₄)) and their ratio, i.e. the relative error errorrel according to Eq. (6). See Sect. 2.7 for definition of the terms.

<table>
<thead>
<tr>
<th>Sample point</th>
<th>drift span [µmol mol⁻¹]</th>
<th>Q₁(Δoff₁) [µmol mol⁻¹]</th>
<th>Q₄(Δoff₄) [µmol mol⁻¹]</th>
<th>errorrel</th>
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</thead>
<tbody>
<tr>
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<td>9.4</td>
<td>-0.46</td>
<td>0.41</td>
<td>0.09</td>
</tr>
<tr>
<td>M6</td>
<td>10.4</td>
<td>-0.86</td>
<td>0.72</td>
<td>0.15</td>
</tr>
<tr>
<td>M7</td>
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<td>0.47</td>
<td>0.12</td>
</tr>
<tr>
<td>M8</td>
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<td>-0.7</td>
<td>0.49</td>
<td>0.16</td>
</tr>
<tr>
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<td>0.41</td>
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<tr>
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<td>M14</td>
<td>12.4</td>
<td>-0.58</td>
<td>0.56</td>
<td>0.09</td>
</tr>
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</table>

Fig. 1. Sampling locations for sub-canopy CO₂ concentration at a 2.25 m height. Distances between sampling points are given in meters. M-numbers are used for reference in the text. Topography is shown by isolines with an equidistance of 0.2 m relative to 750 m a.s.l.
Fig. 2. Hypothetical concentration time series $c_i(t)$ and $c_j(t)$ with time $t \in [1,10] \ (a, c)$, and corresponding frequency and density distributions of concentration differences $c_i(t) - c_j(t) \ (b, d)$ for mixed conditions (a, b) and for non mixed conditions (c, d). Regarding the density distributions (b and d), the histogram bars indicate the frequency for binwidths of 1.0, the solid line is a kernel density estimation generated with the same tools which were used for density estimation of measured concentration data as described in Sect. 2.7.

Fig. 3. Setup of Large Eddy Simulation study. Virtual sensor locations (a), Source-sink distribution (b) and background concentration gradient (c). Grid spacing: 5 m.
Fig. 4. (a) Density distribution of mixing index $MI$ (solid line). Dashed lines at $MI=0.06$ and $MI=0.12$ enclose range for sensible choices of a critical mixing index $MI_c$. (b) Diurnal course of mixing index on 29 June 2008 (solid line) and $MI_c$ (dashed line). $MI$ is representative for the whole sample point field (see Sect. 2.6 for details of the calculation).

Fig. 5. Density distribution of LES modelled concentration differences $c_i(t) - \bar{c}(t)$ of a point measurement $c_i(t)$ relative to the field average concentration $\bar{c}(t)$ for concentration time series $c_1(t), c_2(t), ..., c_{16}(t)$ and $n = 16$ sensor locations 1, 2, ..., 16 (a), and for $c_1(t)$ and $c_{12}(t)$ at sensor locations 1 and 12 (b). Note that the density distributions of $c_1(t) - \bar{c}(t)$ and $c_{12}(t) - \bar{c}(t)$ have a common mode but different mean.
Fig. 6. Calibrated CO$_2$ concentration time series, (a) before bias correction and (b) after bias correction with $M_{I_i} = 0.12$, measured at ten sub-canopy locations M5, M6, …, M14 at a 2.25 m height with a 1 Hz resolution on 29 June 2008.

Fig. 7. Density distribution of concentration differences $c_i(t) - \tilde{c}(t)$ using measured 60-minute concentration timeseries $c_i(t)$ before bias correction (a) and after bias correction (b). Number of sample locations $n = 10$ (M5, M6, …, M14), on 29 June 2008, 12:00–13:00. Legend indicates measurement locations according to Fig. 1.
Fig. 8. Density distribution of concentration differences $c_i(t) - \bar{c}(t)$ using bias corrected measured 60-minute concentration timeseries $c_i(t)$, number of sample locations $n = 10$, for five typical cases over the course of the day on 29 June 2008 with varying stability parameter $\zeta$ (measured at a $36 \text{ m}$ height) and mixing index $MI$ (according to Eq. 3), night time, 01:00–02:00, $\zeta = -0.16$, $MI=0.015$ (a), night-day transition, 07:00–08:00, $\zeta = 0.65$, $MI=0.020$ (b), daytime, 12:00–13:00, $\zeta = -0.27$, $MI=0.218$ (c), day-night transition, 19:00–20:00, $\zeta = 0.06$, $MI=0.010$ (d) and nighttime with katabatic drainage flow, 22:00–23:00, $\zeta = 19.50$, $MI=0.016$ (e). Legend indicates measurement locations according to Fig. 1.

Fig. 9. Time series of the modes of density distributions of concentration differences $c_i(t) - \bar{c}(t)$ (see Fig. 7a for example distributions for one 60-minute time step) for 10 sampling locations on 29 June and 30 June 2008, (a) before filtering with mixing index, and (b) after filtering with mixing index $MI_c=0.13$. Modes from periods which satisfy $MI < MI_c$ are not used during bias correction (grey mask). The last mode at a time with $MI \geq MI_c$ is used instead (solid lines).
Fig. 10. Modes of 60-minute density distributions of concentration differences $c_i(t) - \bar{c}$ minus analyzer drift (mode of 24 h pdf of 60-minute modes subtracted daily) versus mixing index $MI$. Grey points indicate measurements, the solid line marks the 25% and 75% quantiles for mixing index binwidths of 0.025 with the circles centered at each bin. Dashed line at $MI=0.12$ indicates a sensible choice for the critical mixing index $MI_c$. Due to the dependence of natural concentration differences on the mixing index and atmospheric stability, both of which have a distinct daily cycle. Figure 10 illustrates the dependence of concentration offsets on the mixing index $MI$. For low values of $MI$, natural horizontal gradients are large, as a result of horizontal source heterogeneities and potential mixing of a vertical concentration profile with large vertical gradients. For larger values of $MI$, particularly for $MI>MI_c$, offsets are relatively small. Figure 10 displays the dependence of only natural concentration differences on $MI$. The graph shows offsets with instrument bias removed by subtracting a low frequency component (see figure caption for details). The fact that this technique does not perfectly separate...